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Does Composition of Public Expenditure Matter for Economic Growth? Lessons from Sri Lanka

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Abstract: In the context of surging public expenditure and crumbling output growth, the growth effects of public expenditure have provoked an extensive discussion in the economic and political arenas in Sri Lanka. Since 1977, both public expenditure and its composition have changed intensely and largely been accompanied by expansion in size of successive governments. Although it is difficult to determine whether Sri Lanka has reached its optimal size of public expenditure, understanding the growth effects of public expenditure would clearly link policy contributions made by public expenditure in spurring growth in Sri Lanka. The purpose of this study is to examine the growth effects of composition of public expenditure considering full implications of government budget constraints. This study considers public expenditure at a disaggregated level to isolate productive elements of public expenditure from the total. Accordingly, public expenditure on education, health, defence, agriculture and transport and communication are considered. These expenditure items are selected based on their share in total expenditure. This study found that the growth effects of public expenditure vary at disaggregated levels. A major finding showed that public expenditure in education, agriculture, transport and communication sectors is positively and significantly associated with economic growth while defence and health expenditure do not have any significant impact on growth. Given the high magnitude of positive and significant growth effects of public expenditure in the education sector, this study suggests reforming public expenditure in favour of human capital development is paramount to stimulate long-term growth in Sri Lanka.

1. Introduction

In the recent years, the growth effects of size and the nature of public expenditure have emerged as a major issue in economies that are in transition. Public expenditures are less flexible than fiscal revenues, but much more sensitive with regard to business cycles and policy decisions of the government. Government collects revenue through various taxes and allocates to several sectors of the economy to pursue number of objectives. However, allocation of such expenditures is directly and indirectly associated with growth in the

respective economies (Barro, 1990; Tanzi and Zee, 1997; Bayraktar, *et al.* 2015). For instance, the supply of social and physical infrastructure, rule of law, and protection of property rights are assumed to be conducive for growth (Ram, 1986). Over the period, public expenditure policies in both advanced and emerging economies mainly aimed at promoting sustained and equitable economic growth. However, literature highlights growth effects of public expenditure are positive when the size of government is small, but it may become negative as the size gets larger (Grossman and Helpman, 1991).

In the context of surging public expenditure and crumbling output growth, the growth effects of public expenditure have provoked an extensive discussion in the economic and political arenas of Sri Lanka. Since 1977, both public expenditure and its composition have changed intensely and largely been accompanied by expansion in size of successive governments (Figure A2 in Appendix). However, on the empirical front, the existing studies for Sri Lanka (Herath, 2010; Lahirushan, and Gunasekara, 2015; Dilrukshini, 2004) have focused on either the growth effects of total public expenditure or the specific individual expenditure. To the best of our knowledge, a few studies have examined the growth effects of public expenditure at disaggregated levels (Kesavarajah, M and Ravinthirakumaran, N, 2011). This study, therefore, attempts to fill an existing gap in the literature in light of more recent evidence.

The present study departs from existing literature in four key aspects. First, this study uses a time series dataset covering the entire post-liberalization period from 1977 to 2016 and places special focus on the presence of structural breaks in both public expenditure and economic growth series¹. Secondly, public expenditure will be used as opposed to government consumption expenditure. Thirdly, this study will consider public expenditure at a disaggregated level to isolate productive elements of public expenditure from total. Finally, the analytical approach adopted in this study will be different from previous studies.

Although it is difficult to determine whether Sri Lanka has reached its optimal size of public expenditure, understanding the growth effects of public expenditure would clearly link policy contribution made by public expenditure in spurring growth. Therefore, the main research question this study attempts to address is: have growing public expenditure really helped in stimulating economic growth in Sri Lanka. In addressing this question, a fundamental issue then is what components of public expenditure might be conducive or detrimental to economic growth in Sri Lanka? This study aims to provide fresh empirical insights for these research questions. Given the research questions,

the primary objective of this study is to examine the growth effects of components of public expenditure in Sri Lanka. In this study, public expenditure on education, health, defense, agriculture and transport and communication are considered. These expenditure items are selected based on their share in total expenditure.

The rest of this paper is structured as follows. Section two briefly reviews theoretical and empirical literature relating to the relationship between public expenditure and economic growth. Section three highlights the data, econometric models and analytical framework adopted in this study. Section four offers quantitative insights on growth effects of both total and components of public expenditure. Final section summarizes major findings of the study, recommends appropriate policy responses, and suggests avenues for further research.

2. Literature Review

2.1. Theoretical Review

There are three main theories, namely Wagner's law, Keynesian growth theory and endogenous growth theory, which discuss theoretical relationships between public expenditure and economic growth. According to Wagner's law, public expenditure is an endogenous factor driven by growth of national income. It further states that economic activities undertakenby the government upsurges compared to the private sector during economic development (Wagner, 1883). In contrast, Keynesian growth theory (1936) considers public expenditure as an important exogenous variable in determining growth. It argues, given the assumption of price rigidity and possibility of excess capacity, expansion in fiscal policy stimulates growth through growing aggregate demand, which affects technical progress. Although both theories focus on short-run phenomenon of public expenditure, the causality between public expenditure and growth highlighted by these theories is different. According to Keynesian growth theory, causality runs from public expenditure to growth. Wagner's law presents the opposite conclusion. However, several studies including Devarajan, et al. (1996), Afonso and Furceri (2008) and Bose et al. (2007) highlight that the nexus between public expenditure and growth depends on the nature of the expenditure.

The endogenous growth theory highlights important factors that contribute to cross-country differences in both per capita income and growth rates. These factors comprise, investment in human capital (Lucas, 1988), knowledge spillovers, and investment in physical infrastructure. This theory argues that government's policies, including fiscal policy, can affects long-run growth. Although these theories highlight that publicexpenditure affects growth in several channels, empirically, however, each channel leads to varied conclusions.

2.2. Empirical Evidence

Beyond theoretical outpourings, several empirical studies have investigated growth effects of public expenditure in both developed and developing economies. However, findings of these studies have brought inconclusive results. Some studies show that public expenditure has positive impacts on growth while others show detrimental impacts. Studies have also found neutral growth effects of public expenditure. Differences in outcome could be largely due to the nature of data and differences in econometric techniques.

A study conducted by Barro (1990) integrating both developed and developing economies for the period 1960 to 1985 shows that the nexus between public expenditure and growth is weak. Komain and Brahmasrene (2007) focusing on Thailand find a significant positive impact of public expenditure on growth. They also highlight that a unidirectional causality goes from public expenditure to growth. Using cross section data for 71 countries, Cooray (2009) shows that the growth effects of both government size and quality of governance are positive. Castles and Dowrick (1990) use composition of public expenditure and find that social transfers and education expenditure have a positive impact on economic growth. A similar study was conducted by Ranjan and Sharma (2008) for the case of Indian economy and found growth effects of public expenditure to be positive and significant during the period 1950 to 2007.

Tanzi and Zee, (1997) find that public expenditure on infrastructure, human capital, science and technology influences positively on economic growth. Avila and Strauch (2003) show that the expenditure side of budget consistently affected growth in 15 member countries in European Union for the period 1960 to 2001. The study argues that investment expenditure had positive impacts on growth but expenditure on consumption and transfer payments had significant negative impacts on growth. However, Levine and Renelt (1992) show that growth effects of public expenditure are insignificant. Similarly, Devarajan, *et al.* (1996) investigate growth effects of various types of public expenditure and highlighted components of public expenditure matters for growth.

Although many studies establish positive growth effects of public expenditure, studies also showed opposite results. A study conducted by Folster and Henrekson (2001) for Organization for Economic Cooperation and Development (OECD) countries highlightnegative growth effects of public expenditure. They show countries experienced with higher public expenditure registered lower output growth compared with countries registered low public expenditure. Ramayandi (2003) found adverse impacts of increased public expenditure on growth in Indonesia. Conducting a research for a sample of 96 countries, Landau(1983) concluded expenditure on education and defense sectors has a weak impact on growth. He also confirmed that growth effects of total expenditure are negative.

Few studies have examined growth effects of public expenditure in Sri Lanka and results are inconclusive. Herath (2010) covering the period 1959 to 2003 found positive and significant impacts of public expenditure on economic growth. He also confirms positive growth effects of openness. Similar results have been established by Lahirushan and Gunasekara (2015) while confirming bidirectional causality between public expenditure and growth. Meanwhile, Jayawickrame (2004) stresses reduction in government consumption expenditure, transfer payments and investment expenditure, which have adverse impacts on growth. However, Dilrukshini (2004) shows that there is no empirical evidence in support of either Wagner's Law or the Keynesian hypothesis in Sri Lanka.

3. Data, Econometric Models and Analytical Framework

3.1. Data

This study uses time series annual data for the period 1977 to 2016. The data on fiscal variables are mainly based on annual reports of the Central Bank of Sri Lanka (CBSL). Data for other variables have been extracted from three different sources. Growth rates of GDP and gross fixed capital formation are mainly drawn from World Development Indicators. Further, growth rates of population and inflation are drawn from various publications of the Department of Census and Statistics (DCS) of Sri Lanka. The data on human capital variables are drawn from Barro and Lee (2000) data set. Detailed explanation on variables and their sources are given in Table A1 in Appendix.

3.2. Econometric Models

The analysis of this study will be based on the standard growth regression model. The standard growth regression model, which follows seminal contribution of Barro (1990), has been widely used in empirical research on economic growth. This model is based on a conditional convergence equation that relates output growth to initial levels of income, investment, human capital and population growth. However, to achieve the main objective of the present study, we augment standard growth regression model with both fiscal and nonfiscal variables. Accordingly, the basic empirical model is specified as follows.

$$GR_t = \boldsymbol{\beta}_0 + \sum_{j=1}^p \boldsymbol{\beta}_j X_t + \sum_{j=1}^q \boldsymbol{\beta}_j Y_t + \sum_{j=1}^r \boldsymbol{\beta}_j Z_t + \boldsymbol{\varepsilon}_t$$
(1)

Where, t is the year index, GR is the growth rate of real GDP, β_0 and β_j are coefficients in regression model. ε_j is the error term assumed to be white noise process. Growth rate of real GDP is dependent variable and explanatory variables are classified into three groups: X, Y and Z. Group X consists of conditioning variables that are commonly appears in growth regression model, while group Y includes non-fiscal control variables that are generally included in empirical literature on economic growth. Fiscal variables that we are interested in this study are included in group Z.

Consistent with advocates of growth theory, we include gross investment, growth rate of population and human capital as conditioning variables. Although inclusion of different control variables results to different outcomes, the selections of control variables in this study are based on growth literature. As many countries registered higher growth consistent with export-led strategies, we include trade openness as the control variable. Following Easterly and Rebelo (1993), we include broad money and inflation to capture growth effects of monetary policy and macroeconomic instability. Further, lagged value of economic growth is included to capture growth inertia factors. We also introduce a dummy variable to examine the impacts of civil war on growth.

As Sri Lanka is confronted with fiscal constraints over the periods, we considered government budget constraints to eliminate coefficient bias resulting from their omission (Bose *et al.* 2007). Accordingly, we include public expenditures, tax revenues and fiscal balance to growth regression model. We also exclude some expenditure items to avoid any multicollinearity problem. As there is a time lag between the execution of public expenditure and its propagation on economy, we consider lagged value of public expenditure in the model.

Since we aim to examine growth effects of both total and components of public expenditure, we first estimate growth effects of total public expenditure as specified in equation 2 and then we jointly include five components of public expenditure as given in equation 3 to isolate productive elements of public expenditure from total.

$$GR_{t} = \boldsymbol{\beta}_{0} + \sum_{j=1}^{p} \boldsymbol{\beta}_{j} X_{t} + \sum_{j=1}^{p} \boldsymbol{\beta}_{j} Y_{t} + \boldsymbol{\beta}_{1} E X P_{t} + \boldsymbol{\beta}_{3} T R_{t} + \boldsymbol{\beta}_{4} F B_{t} + \boldsymbol{\beta}_{5} G R_{t-1} + \boldsymbol{\beta}_{6} D war_{t} + \boldsymbol{\varepsilon}_{t}$$

$$(2)$$

$$GR_{t} = \boldsymbol{\beta}_{0} + \sum_{j=1}^{p} \boldsymbol{\beta}_{j} X_{t} + \sum_{j=1}^{p} \boldsymbol{\beta}_{j} Y_{t} + \boldsymbol{\beta}_{1} EDU_{t} + \boldsymbol{\beta}_{1} HEL_{t} + \boldsymbol{\beta}_{1} TRC_{t} + \boldsymbol{\beta}_{1} AG_{t} + \boldsymbol{\beta}_{1} DEF_{t} + \boldsymbol{\beta}_{3} TR_{t} + \boldsymbol{\beta}_{4} FB_{t} + \boldsymbol{\beta}_{5} GR_{t-1} + \boldsymbol{\beta}_{6} Dwar_{t} + \boldsymbol{\varepsilon}_{t}$$

$$(3)$$

3.3. Analytical Framework

The analysis of this study will be conducted in five stages. Given that our sample involves the period around the end of the civil war in 2009 and the emergence of global financial crisis in 2008, we will first investigate structural changes in both economic growth and public expenditure series. The global maximiser test introduced by Bai and Perron (2003) will be used to determine any structural breaks.

As many time series variables contain unit root, in the second stage, we will examine stationary properties of all the variables using Augmented Dickey Fuller (ADF), Phillips Perron (PP) and Kwiatkowski-Phillips-Schmidt-Shin (KPSS) tests. We will use Akaike Information Criteria (AIC) with default lag order for selecting order of augmentation in ADF regression. Further, default settings of Bartlett Kernel and Newey-West Bandwidth will be used for KPSS tests.

Although Engel-Granger (1987) and Johenson and Juselious (1990) methods are widely used to examine co-integration among variables, these techniques are not reliable for small sample size and may not provide co-integration even when all variables are integrated with order I(1) (Kremers at el, 1992). Therefore, in the third stage, we will use the Autoregressive Distributed Lag (ARDL) approach introduced by Pesaran and Pesaran (1997) to examine long-run dynamics among variables. This approach can be applied even when variables are I(0) or I(1) or mixture of both. It also considers adequate number of lags to capture data generating process. Given the scope of this study, the ARDL model is specified as follows.

$$\Delta GR_{t} = \beta_{0} + \beta_{1}GR_{t-1} + \beta_{2}X_{t-1} + \beta_{3}Y_{t-1} + \beta_{4}Z_{t-1} + \sum_{j=1}^{p}\lambda_{j}\Delta GR_{t-j} + \sum_{j=1}^{p}\Psi_{j}\Delta X_{t-j} + \sum_{j=1}^{p}\theta_{j}\Delta Y_{t-j} + \sum_{j=1}^{p}\Pi_{j}\Delta Z_{t-j}$$
(4)

Equation 4 estimates (p+1)k number of regressions to obtain optimal lag length for each variable in the model. Where, p is the optimal lag to be used and k is the number of variables. Δ is the first difference operator. As Schwarz Information Criteria (SIC) is widely known as parsimonious model and selects smallest possible lag length, we will use SIC to select optimal lag length in this study. If variables are found to be co-integrated, we will estimate ARDL Error Correction Model (ECM) to investigate short-run dynamics among variables. The error correction version of modified ARDL model is specified as follows.

$$\Delta GR_{t} = \sum_{b=1}^{p_{1}} \lambda_{b} \Delta GR_{t-b} + \sum_{j=1}^{p_{2}} \Psi_{j} \Delta X_{t-j} + \sum_{s=1}^{p_{3}} \theta_{s} \Delta Y_{t-s} + \sum_{i=1}^{p_{4}} \Pi_{i} \Delta Z_{t-i} + \gamma E C_{t-1} + \varepsilon_{t}$$
(5)

The coefficient (γ) of EC, which shows the speed of adjustment on a yearly basis to long-run equilibrium after a short-run deviation, should be negative and significant. Finally, we will conduct appropriate diagnostic and stability tests to ensure goodness of fit of estimated ARDL model. The stability testing will be conducted using both Cumulative Sum (CUSUM) and Cumulative Sum of Squares (CUSUMSQ) methods suggested by Pesaran and Pesaran (1997).

4. Results and Discussion

4.1. Testing for Structural Breaks

The structural break tests based on global maximiser test for both public expenditure and economic growth confirmed absence of any structural break in both series (see Table B1 and B2 in Appendix B). Therefore, we proceed with the next step to examine order of integration of the variables.

4.2. Unit Root Test

It is noted that except the human capital variable, which exhibits an upward trend, all other variables do not exhibit a clear trend (see Figure A1 in Appendix A). Hence, we included only an 'intercept' component in ADF test. It is confirmed that twelve variables are stationary in levels (see Table 1). However,

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Variables	ADF Level	+ 1 est First Difference	Level Level	1 est First Difference	K ^I .X.	3 1 est First Difference	Order of Integration
	-4.7984***	-9.1672***	-4.7945***	-25.6584***	0.0871	0.5000**	I(0)
X Variables							
inv	-3.7531**	-5.0329***	-3.9757**	-5.7959***	0.0770	0.1242^{**}	I(0)
pg	-5.6943***	-6.7098***	-6.7001***	-27.1822***	0.1529^{**}	0.5000 **	I(0)
hucp	-1.7919	-6.7334***	-1.6342	-7.5965***	0.1663^{**}	0.2490	I(1)
Y Variables							
ddo	-3.2233***	-8.0797***	-3.4230**	-8.3629***	0.1718^{**}	0.2818	I(0)
m2	-1.9517	-6.1518***	-2.4967	-6.1777***	0.0853	0.1337 **	I(1)
inf	-5.6249***	-6.0579***	-5.6258***	-17.6133***	0.0629	0.2825^{**}	I(0)
Z Variables							
pexp	-5.9531 * * *	-6.5428***	-10.3866***	-29.0965***	0.5000^{**}	0.2721	I(0)
fb	-6.4543***	-10.8657	-6.3894***	-13.3118*	0.0659	0.1223^{**}	I(0)
trev	-3.7997***	-10.3300 ***	-5.4604***	-10.7585*	0.1202^{***}	0.5000	I(0)
edu	-3.2261*	-7.7934***	-3.2168***	-8.2004***	0.1552^{**}	0.0672	I(0)
hth	-3.8563**	-7.1914***	-3.9318**	-10.3078 ***	0.0608	0.6894^{**}	I(0)
trc	-3.7958***	-9.3847***	-3.7665***	-9.8155***	0.0919	0.1669 **	I(0)
agr	-6.8661***	-3.5260**	-2.9376	-7.5798***	0.1793^{**}	0.0613	I(0)
def	-1.5893	-6.3729***	-1.4553	-6.4137***	0.1839^{**}	0.2627	I(1)

Table 1: Unit Root Tests – Levels and First Difference

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Source: Author's Estimation

it is found that three variables are non-stationary since the null hypothesis of series being stationary cannot be rejected in both ADF and PP tests at 5 percent level of significance. However, all three tests confirmed that the remaining variables are integrated with order one. Since variables used in this study have a mixed order of integration, we proceed with estimating ARDL models for co-integration testing.

4.2. Estimation of ARDL Models

4.2.1. ARDL Model for Growth Effects of Total Public Expenditure

We estimated 20 ARDL models to examine growth effects of total public expenditure. The graphical representations of these 20 models are given in Figure 1. Based on the SIC, the ARDL model $(1\ 0\ 1\ 0\ 1\ 1\ 0\ 0\ 1\ 1)$ was selected as the best model among them. Further, we conducted a stability test for estimated ARDL model $(1\ 0\ 1\ 0\ 1\ 1\ 0\ 0\ 1\ 1)$. Since both CUSUM and CUSUMSQ statistics prevailed within the critical bounds of 5 percent level of significance, the null hypothesis of estimated coefficients of ARDL model is stable cannot be rejected (see Figure 2 and 3). This confirms that estimated ARDL model is stable.



Figure 1: Model Selection Summary - Schwarz Information Criteria (top 20 models)

Note: The ordering of the variables is GR, PEXP, M2, OPP, INF, FB, TREV, HUCP, INV, PG.

Detailed descriptions of the variables are given in Appendix Table A1.



Figure 2: CUSUM Test







Source: Author's Estimation

As estimated ARDL model is stable, we conducted ARDL-Bounds test to identify co-integration among variables in the model. The results of Bounds

test are presented in Table 3. As calculated F-statistic is higher than upper bound critical values, the null hypothesis of no long-run (equilibrating) relationship can be rejected in all three significance level. This confirms the existence of a long-run relationship irrespective of order of integration of variables.

Test Statistic	Value	I(0)): Lower Boi	und	I(1): Upper Bo	und		
	Asymptotic: n =1000								
F-Statistic	6.78265	Significance Level			Si	ignificance Le	evel		
		10%	5%	1%	10%	5%	1%		
Κ	9	1.8	2.04	2.5	2.8	2.08	3.68		

Table 3: ARDL Bound Tests for Long-run Relationship

Source: Author's Estimation

Note: The critical value bounds are computed by stochastic simulations using 1000 replications. K represents the number of variables that we have included in the ARDL bound test.

4.2.2. ARDL Model for Growth Effects of Components of Public Expenditure

We estimated 20 ARDL models based on SIC to examine growth effects of components of public expenditure (see Figure 4). However, we selected ARDL (1, 0, 0, 1, 1, 0, 0, 1, 1, 0, 1, 0, 1) model as the best model among all other models. Moreover, the stability of the estimated ARDL model was investigated using both CUSUM and CUSUMSQ methods. Since both CUSUM and CUSUMSQ statistics prevailed within the critical bounds of 5 percent level of significance, the null hypothesis of estimated coefficients of ARDL model is stable cannot be rejected. This confirms that the estimated ARDL model is stable. As estimated ARDL model is stable, we applied ARDL Bounds tests to examine the long-run relationship among variables. Accordingly, it is found that the F-statistic value of 10.48238 exceeds the critical value of upper bounds in all three significance levels (see Table 4). Therefore, the null hypothesis of no long-run relationship can be rejected in all three significance levels and it can be concluded that all the variables move together in the long-run irrespective of their order of integration.



Figure 4: Model Selection Summary - Schwarz Criteria (top 20 models)

Note: The ordering of the variables is GR, PEXP, M2, OPP, INF, FB, TREV, HUCP, INV, PG.



Source: Author's Calculation

Source: Author's Estimation



Figure 6: CUSUM of Squares Test

Source: Author's Estimation

Table 4: ARDL Bounds Tests for Long-run Relationship

Test Statistic	Value	I(0	I(0): Lower Bound		I(1): Upper Bo	und
			ic: n =1000	00			
F-Statistic	10.48238	Significance Level			Significance Level		
		10%	5%	1%	10%	5%	1%
К	13	1.76	1.98	2.41	2.77	3.04	3.61

Note: The critical value bounds are computed by stochastic simulations using 1000 replications.

K represents the number of variables included in the model.

Source: Author's Estimation

4.3. The Long-Run Impacts of Public Expenditure on Economic Growth

4.3.1. Total Public Expenditure

As estimated models in section 4.2.1 and 4.2.2 established a co-integration among variables, in the next step, we proceed with estimating long-run coefficients using respective ARDL specifications in two different settings. The first set of

estimations, which includes total public expenditure along with other control variables, is presented in the second column of Table 5. The second set of estimation, which includes components of public expenditure, is given in the last column of Table 5.

As per model 1, total public expenditure has negative and significant impacts on growth. Holding other variables of the model constant, a one percent increase in public expenditure as a percentage of GDP is associated with approximately 0.18 percent decrease in long-term growth. This result is consistent with previous studies (Devarajan *et al.* 1996; Folster and Henrekson 2001), which argue that public expenditure has negative impacts on growth owing to inefficient and unproductive nature of such investments. However, the result contradicts with a previous study conducted by Herath (2010) who found positive growth effects of public expenditure for Sri Lanka. This conflicting result could be due to the inclusion of most recent data that captures recent changes in the economic structure and the usage of advanced econometrics techniques. However, given the increased public expenditure, examining the growth effects of public expenditure at disaggregated level is imperative for Sri Lanka. Therefore, a detailed analysis is conducted in the subsequent section.

Dependent	Variable: Growth rates of rea	l GDP
Variable	Results with Total Public Expenditure (Model 1) ARDL Model: (1, 0, 1, 0, 1, 1, 0, 0, 1, 1)	Results with Components of Total Public Expenditure (Model 2) ARDL Model: (1, 0, 0, 1, 1, 0, 0, 1, 1, 1, 0, 1, 0, 1)
X Variables		
Investment (% of GDP)	0.5863***	0.4264**
	(4.9316)	(2.1360)
Population Growth Rate	0.3783	0.7487
-	(0.3021)	(1.1874)
Human Capital	0.0505	0.1118***
-	(1.0565)	(2.9504)
Y Variables		
Trade Openness ((% of GDP)	6.2649	0.2680
	(1.6135)	(0.6501)

 Table 5: Estimated Long-Run Coefficients of ARDL Model with Budget Constraint

contd. table 5

Dependent Variable: Growth rates of real GDP					
Variable	Results with Total Public Expenditure (Model 1) ARDL Model: (1, 0, 1, 0, 1, 1, 0, 0, 1, 1)	Results with Components of Total Public Expenditure (Model 2) ARDL Model: (1, 0, 0, 1, 1, 0, 0, 1, 1, 1, 0, 1, 0, 1)			
Money Supply (M2) (% of GDP)	0.1447	0.0406			
	(0.8153)	(0.3222)			
Inflation Rate	0.1008	0.0314			
	(1.3230)	(1.5883)			
Growth Rates of GDP (-1)	-0.9485***	-1.0527***			
	(-5.6317)	(-8.1275)			
DWAR	-2.9306***	-0.7769			
	(-2.6083)	(-0.7975)			
Z Variables					
Total Expenditure (-1) (% of GDP)	-0.1797***				
	(-3.6259)				
Tax Revenue (% of GDP)	-0.4571	-0.7353***			
	(-0.7487)	(-2.6989)			
Fiscal Balance (% of GDP)	0.2204**	0.2617			
	(2.3135)	(1.4411)			
Education Expenditure (-1)(% of G	DP)	5.8003***			
		(3.3396)			
Health Expenditure (-1)(% of GDP)		0.7249			
		(0.3695)			
Transport and Communication Expenditure (-1) (% of GDP)		0.7831*			
		(1.9242)			
Agriculture Expenditure (-1) (% of GDP)		1.2721***			
		(3.3216)			
Defense Expenditure (-1) (% of GDP)		-0.2545			
		(-0.6548)			
Constant	-10.1814	-14.8284			
	(-0.7963)	(-1.5168)			
No. of Observations	40	40			

Notes: ***indicates significance at 1%, ** indicates significance at 5% and * indicates significance at 10% levels.

t-statistics are given in parenthesis. (-1) indicates variables lagged one period

Source: Author's Estimation

4.3.2. Components of Public Expenditure

As evident from model 2, growth effects of total expenditure on education are found to be positive and statistically significant (see Table 5). The magnitude of growth effects of education expenditure is substantial. This result is consistent with new growth literature which strongly validates the argument of education expenditure as an essential factor in determining growth. Holding other variables constant, one percent increase in expenditure in education as a percentage of GDP is associated with an increase in average growth rate by 5.8 percent. Previous studies by Castles and Dowrick (1990) also established similar results. The positive impacts of education expenditure could be due to strong spillover effects of investment in the education sector in raising the productivity of both human and physical capital. However, our results contradict with previous studies conducted by Barro and Sala-i-Martin (1992), Landau (1983) and Devarajan *et al.* (1996) that established insignificant impacts of education expenditure.

Health expenditure is considered as an investment in human capital that influences growth. However, this study found that health expenditure has a positive but insignificant impact on growth. Our finding contradicts with previous studies of Cole and Neumayer (2006) who showed positive and significant growth effects of health expenditure. Inclusion of budget constraints, which is absent in previous studies, and is limited to only Sri Lankaneconomy could partly explain reasons for different outcomes. To our knowledge, studies on growth effects of health expenditure are almost non-existent in the literature on Sri Lanka and therefore this study provides fresh empirical evidence.

Meanwhile, this study found growth effects of transport and communication expenditure are positive and significant. The result is consistent with previous studies conducted by Fedderke *et al.* (2006), who confirmed positive impacts of infrastructure on growth in South Africa. The present study also stresses that expenditure on transport and communication sector is crucial for growth. Furthermore, this study clearly establishes positive and significant growth effects of agriculture expenditure. Therefore, allocating funds towards agriculture sector would not only enhance growth but could also improve the wellbeing of rural population. Meanwhile, coefficient on defence expenditure is negative but statistically insignificant. This is consistent with previous studies such as Landau (1983) and Deger and Smith (1983). The diversion of expenditure towards unproductive sectors including defense results in a reduction of public savings and investments and thereby undermines growth. Although this study found adverse impacts of defence expenditure, some positive impacts are also justified in literature (see Frederiksen and Looney, 1983; Ram, 1986).

Looking at other variables that are considered in this study, while model 1 does not provide any evidence on growth effects of tax revenue, model 2 highlights negative and significant effects. The negative impacts of tax revenue could be identified from different channels. Increased taxation discourages both domestic and international investment and thereby adversely affects long-term growth. It also adversely affects investment in human capital and entrepreneurial activities. However, investigating growth effects of components of tax revenue is beyond the scope of this study. The budget surplus has significant positive impacts on growth in both models. Holding other variables constant, one percent increase in fiscal surplus as a share of GDP is associated with an increase in annual growth rate by an average of 0.22 and 0.26 percent in model 1 and 2 respectively. This highlights that controlling fiscal deficit is vital to enhance long-term growth.

Looking at growth effects of non-fiscal variables, the study found that investment played a crucial role in stimulating growth, which is consistent with neoclassical growth theory. Several previous studies including a study by Levine and Renelt (1992) also acknowledge investment as important determinants of growth. Holding other variables constant, one percent increases in total investment as a percentage of GDP is associated with an increase in growth rate by 0.6 percent in model 1 and 0.4 percent in model 2. This necessitates the creation of a conducive environment in Sri Lanka to attract both domestic and global investment. Human capital is found to have a positive impact on growth in both models though significance can be observed only in model 2. While inflation, population growth, money supply and trade openness accord well with theoretical predictions, they are turn out to be insignificant in both models. It is evident that growth can be accounted for by its own innovations. The study also confirmed that civil war had negatively affected growth in both models though the significance could be identified only in model 1.

4.4. Estimation of ARDL Error Correction Model (ECM)

As the Bounds test established a long-run co-integration among variables, we will use ECM of ARDL to estimate short-run dynamics among the variables. The error correction term is negative and significant in both models. This shows a high level of speed of adjustment to long-run equilibrium following a short-run shock (see Table 6). About 34.41 percent and 52.70 percent of disequilibrium

from previous year's short-run shocks converges back to long-run equilibrium each year in model 1 and 2.

Dej	pendent Variable: Growth rates of rea	l GDP
Variable	Results with Total Public Expenditure[ARDL (1, 0, 1, 0, 1, 1, 0, 0, 1, 1)]	Results with Components of Total Public Expenditure [ARDL (1, 0, 0, 1, 1, 0, 0, 1, 1, 1, 0, 1, 0, 1)]
X Variables		
D(INV)	0.3441***	0.5841***
	(3.9959)	(5.6434)
D(PG)	0.3783	0.0314
	(0.8623)	(1.5883)
D(HUCP)	0.1039	
	(1.2609)	
Y Variables		
D(M2)	-0.6560***	
	(5.3981)	
D(INF)	-0.0939***	-0.6632*
	(-2.9265)	(-1.9409)
WRDUM	-2.9306***	-7.7612***
	(-8.2899)	(-16.4002)
Z Variables		
D(PEXP(-1))	0.0020	
	(0.0839)	
D(FB)	0.6889***	0.5593***
	(7.6513)	(10.5940)
D(EDU)		3.7185***
		(7.6721)
D(HTH)		0.0049
		(0.0772)
D(AGR)		0.2076**
		(1.9636)
ECM (-1)	-0.3441***	-0.5270***
	(-3.9959)	(-8.5970)
R-squared	0.8307	0.9274
Adj. R2	0.7989	0.9079
Durbin-Watson stat	1.9792	2.04265

Table 6: ARDL Error Correction Model of the Growth Equation

Notes: ***indicates significance at 1%, ** indicates significance at 5% and * indicates significance at 10% levels. Source: Author's Calculation

Although signs of short-run coefficients of total public expenditure are in line with Keynesian theoretical outpourings, its impact on growth is insignificant. It is found that total expenditure on education and agriculture, has a similar impact in the short-run too but with different magnitudes. However, the transport and communication expenditure that exhibited positive and significant impacts on long-run growth do not provide any evidence in the short-run. Growth effects of health expenditure is insignificant in the short-run as well. Similar to long-term results, growth effects of fiscal balance are positive though impact is significant only in model 1. However, the degree of impacts is much lower in the short-run compared to the long-run. This result strengthens the argument in favour of rapidly implementing effective policies for deficit reduction in Sri Lanka.

Although growth effects of inflation are insignificant in the long-run, they have a negative and significant impact in the short-run. The main policy message is that reducing inflation by 1 percent could raise growth by 0.66 percent. This provides strong evidence to support the view of low inflation to stimulate shortterm growth. Moreover, investigating channels through which inflation affects growth is essential, but it is beyond the scope of this study. Growth effects of investment exhibit positive and significant results in both models, which is consistent with long-run results. Although long-run growth effects of money supply are consistent with theoretical prediction, the study shows that money supply has significantly affected growth in the short-run. It is also evident that the civil war undermined economic growth in the short-run as well. The shortrun impacts of all other variables are insignificant though they had expected sign in both models.

5. Conclusion

This study provides new empirical understanding on growth effects of public expenditure at both aggregate and disaggregate levels for Sri Lanka considering the full implication of government budget constraints. The present study found that the growth effects of public expenditure vary at the disaggregated level. The growth effects of total expenditure on education and transport and communication were found to be positive and statistically significant, while health, agriculture and defense expenditure were found to be statistically insignificant. Looking at growth effects of non-fiscal variables, the study found that investment and human capital played a crucial role in stimulating growth in Sri Lanka. Given the positive and significant growth effects of expenditure on education, agriculture, transport and communication and negative but insignificant impacts of defense expenditure, this study suggests that reforming public expenditure in particular in favor of human capital development could stimulate long-term growth in Sri Lanka. As it is found that the tax revenue has adversely affected economic growth, the present study recommends that controlling possible detrimental impacts of taxation is vital to enhance growth. Given the positive growth effects of fiscal surplus, this study recommends that controlling fiscal deficit is paramount for sustainable growth. Based on the findings, this study proposes several lines of further investigation and some possible extensions to the research content of this study. As growth effects of public expenditure at disaggregated levels have shown mixed results, in-depth studies focusing on disaggregated components even within these categories should be conducted. Although the present study showed valuable insights on growth effects of tax revenue, it is necessary to address potential growth implications arising from taxation at disaggregated levels.

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Note

1. Sri Lankan economy was liberalised in 1977. The study period covered in this study contains two important break periods. First, the civil war started in 1983 ended in 2009. And second, the global financial crisis which emerged in 2008.

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Appendices

A.1 Information on Data

Figure A1: Behaviour of the Variables

Note: The detailed descriptions of the variables are given in Table A1

Variable	Definition of Variables	Data Source
	GDP Data	
gr	Growth rate of GDP	World Bank Development Indicators
	Conditioning Variables (X)	
inv	Gross fixed capital formation (% of GDP)	World Bank Development Indicators
pg	Population growth rate	DCS
hucp	Human capital (primary, secondary and tertiary school enrolment ratio)	Barro and Lee (1994)
	Control Variables (Y)	
opp	Exports and imports share of GDP (trade openness)	CBSL
m2	Broad Money (M2) (% of GDP)	CBSL
inf	Inflation rate as the percentage change of CPI	DCS
gr(-1)	One period lag of growth rate of GDP	CBSL
wrdum	Dummy variable to capture civil war (1 for 1983- 2009 and 0 for otherwise)	
	Fiscal Variables (7)	
nevn	Public expenditure (% of GDP)	CBSL
pexp(-1)	One period lag of public expenditure (% of GDP)	CBSL
rev	Government revenue (% of GDP)	CBSL
fb	Fiscal Balance (surplus/deficit) (% of GDP)	CBSL
10	Components of Government Expenditure	ODOL
edu	Government expenditure in education (% of GDP) CBSL
hth	Government expenditure in health expenditure (% of GDP)	CBSL
trc	Government expenditure in Transport and Communication (% of GDP)	CBSL
agr	Government expenditure in Agriculture and Irrigat	ion
-	(% of GDP)	CBSL
def	Government expenditure in Defense (% of GDP)	CBSL
	Components of Government Revenue	
trev	Tax revenue (% of GDP)	CBSL

Table A1: Definition of the Variables and Data Sources

Note: CBSL-Central Bank of Sri Lanka, DCS-Department of Census and Statistics of Sri Lanka

A1. Construction of Human Capital Variable

Literatures including Easterly and Rebelo (1993) and Barro and Sala-i-Martin (1999) uses enrolment rates for human capital variable. However, in this study, we follow Bose *et al.* (2007) to construct the initial human capital variable. In this study, the human capital variable is a weighted sum of initial enrolment ratios in primary, secondary and higher education. The weights are 1 for primary school, 2 for secondary school and 3 for higher education. Weights are approximations to the relative values of three types of education. The construction of the human capital variable is based on the Barro and Lee (2000) data source.

Variable	Obs	Mean	Median	Min	Max	Std. Dev.
gr	40	5.1089	5.0999	-1.5454	9.1446	1.9028
pg	40	1.0978	1.1750	-2.2400	2.3500	0.8036
opp	40	0.5782	0.5958	0.3638	0.7741	0.1192
hucp	40	8.8883	9.3500	6.7200	10.3100	1.3528
inv	40	24.7176	24.7965	13.8298	31.3331	3.2970
inf	40	10.2500	9.6000	1.2000	26.1000	5.6426
m2	40	31.1175	30.9000	23.9000	40.7000	2.8616
pexp	40	27.6475	27.1000	17.3000	42.7000	6.1954
fb	40	-8.6025	-8.1000	-19.2000	-4.5000	2.7945
trev	40	15.5325	15.0500	10.1000	24.2000	3.0780
agr	40	2.1499	1.2802	0.6945	10.3932	2.2363
def	40	2.4672	2.4729	0.4897	5.2692	1.2292
edu	40	2.4252	2.4886	1.5597	3.2320	0.4166
hlth	40	1.5174	1.5370	1.1337	2.0169	0.2395
trc	40	2.4758	2.4027	0.7229	5.2158	0.8352

Table A2: Descriptive Statistics of the Variables



Note: The detailed descriptions of the variables are given in Table A1 in Appendix

Figure A2: Trends in Public Expenditure: Disaggregated Analysis

Source: Central Bank of Sri Lanka

Lag	LogL	LR	FPE	AIC	SIC	HQ
0	-659.7271	NA	59.4367	35.3014	35.7755	35.4701
1	-359.8285	410.3875*	0.0062	25.8857	31.5742*	27.9096
2	-183.9458	138.8548	0.0022*	22.9972*	33.9000	26.8763*

Table A3: Selection of Optimal Lag Length

Note: * indicates lag order selected by the criterion, LR: sequential modified LR test statistic, FPE: Final prediction error, AIC: Akaike information criterion, SIC: Schwarz information criterion, HQ: Hannan-Quinn information criterion (each test at 5% significance level)

Source: Author's Calculation

B.1. Testing for Structural Breaks in Public Expenditure and Economic Growth

As our sample involves the time period around the financial crisis of 2008 and the end of the civil war in 2009, we examined structural breaks in public expenditure (PEXP) and conomic growth (GR). This study uses a global maximiser test introduced by Bai and Perron (2003) to determine the years of any structural breaks.

Table B1: Identification of Structural Break for Public Expenditure

Dependent Variable: PEXP Method: Least Squares with Breaks Sample (adjusted): 1981 2016 Break type: Bai-Perron tests of L+1 vs. L sequentially determined breaks Break selection: Trimming 0.15, Max. breaks 5, Sig. level 0.05 No breakpoints selected

Variable	Coefficient	Std. Error	t-Statistic	Prob.
PEXP(-1)	0.560567	0.155598	3.602653	0.0011
PEXP(-2)	0.119634	0.174111	0.687113	0.4971
PEXP(-3)	-0.002657	0.149258	-0.017801	0.9859
PEXP(-4)	0.233271	0.108040	2.159114	0.0387
Non-Breaking Variable	28			
С	1.446762	1.785148	0.810444	0.4239
R-squared	0.877580	Mean depende	ent var	26.69167
Adjusted R-squared	0.861784	S.D. dependen	ıt var	5.219654
S.E. of regression	1.940532	Akaike info cr	iterion	4.292048
Sum squared resid	116.7356	Schwarz criter	ion	4.511981
Log likelihood	-72.25686	Hannan-Quin	n criter.	4.368810
F-statistic	55.55670	Durbin-Watso	on stat	1.892250
Prob(F-statistic)	0.000000			

Source: Author's Calculation

Table B2: Identification of Structural Break for Economic Growth Rate

Dependent Variable: GR

Method: Least Squares with Breaks

Sample (adjusted): 1981 2016

Break type: Bai-Perron tests of L+1 vs. L sequentially determined breaks

Break selection: Trimming 0.15, Max. breaks 5, Sig. level 0.05

No breakpoints selected

Variable	Coefficient	Std. Error	t-Statistic	Prob.
GR(-1)	0.225703	0.179749	1.255657	0.2186
GR(-2)	0.043694	0.183342	0.238320	0.8132
GR(-3)	-0.031467	0.183389	-0.171587	0.8649
GR(-4)	-0.017750	0.180679	-0.098240	0.9224
Non-Breaking Variable	28			
С	3.921429	1.537254	2.550932	0.0159
R-squared	0.057263	Mean depende	ent var	5.037537
Adjusted R-squared	-0.064381	S.D. dependen	t var	1.989329
S.E. of regression	2.052367	Akaike info cr	iterion	4.404111
Sum squared resid	130.5786	Schwarz criter	ion	4.624044
Log likelihood	-74.27400	Hannan-Quini	n criter.	4.480874
F-statistic	0.470743	Durbin-Watso	n stat	1.998211
Prob(F-statistic)	0.756761			

Source: Author's Calculation